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# RESEARCH ARTICLE





# Metabolomic age and risk of 50 chronic diseases in community-dwelling adults: A prospective cohort study

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### Abstract

It is unclear how metabolomic age is associated with the risk of a wide range of chronic diseases. Our analysis included 110,692 participants (training: n=27,673; testing: n=27,673; validating: n=55,346) aged 39–71 years at baseline (2006–2010) from the UK Biobank. Incident chronic diseases were identified using inpatient records, or death registers until January 2021. Predicted metabolomic age was trained and tested based on 168 metabolomics. Metabolomic age was linked to the risk of 50 diseases in the validation dataset. The median follow-up duration for individual diseases ranged from 11.2 years to 11.9 years. After controlling for false discovery rate, chronological age-adjusted age gap (CAAG) was significantly associated with the incidence of 25 out of 50 chronic diseases. After adjustment for full covariates, associations with 15 chronic diseases remained significant. Greater CAAG was associated with increased risk of eight cardiometabolic disorders (including cardiovascular diseases and diabetes), some cancers, alcohol use disorder, chronic obstructive pulmonary disease, chronic kidney disease, chronic liver disease and age-related macular degeneration. The association between CAAG and risk of peripheral vascular disease, other cardiac

Abbreviations: BMI, body mass index; CAAG, chronological age-adjusted age gap; CHD, coronary heart disease; CI, confidence interval; CKD, chronic kidney disease; CMD, cardio metabolic disorder; COPD, chronic obstructive pulmonary disease; CVD, cardiovascular disease; GRS, genetic risk score; HR, hazard ratio.

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diseases, fracture, cataract and thyroid disorder was stronger among individuals with unhealthy diet than in those with healthy diet. The association between CAAG and risk of some conditions was stronger in younger individuals, those with metabolic disorders or low education. Metabolomic age plays an important role in the development of multiple chronic diseases. Healthy diet and high education may mitigate the risk for some chronic diseases due to metabolomic age acceleration.

### KEYWORDS

cardiometabolic disorder, chronic kidney disease, chronic obstructive pulmonary disease, liver disease, metabolomic age, moderation analysis, oesophageal cancer

### 1 | INTRODUCTION

Chronological age is one of the most important risk factors for chronic disorders and mortality (Beard et al., 2016; López-Otín et al., 2013). People with the same chronological age may differ in biological health, and biological ageing may be a better predictor for health (Horvath, 2013; Jylhävä et al., 2017). Biological age has been developed based on omics including genomics, telomere length, transcriptomics and proteomics and it has been shown to be a strong predictor of chronic diseases and mortality (Jylhävä et al., 2017; Kuo et al., 2022). In the global ageing population, it is important to understand biological age for the prevention of chronic diseases and promoting healthy ageing.

Metabolism plays an important role not only in the development of metabolic disorders but also in cardiovascular disease (CVD), neurogenerative disorders, musculoskeletal disorders, mental disorders, respiratory conditions and cancers (Amorim et al., 2022; Eckel et al., 2018; Lumsden et al., 2022). Metabolomic state added predictive value even over established clinical variables in the development of multiple chronic diseases (Buergel et al., 2022). A cohort study from Germany (discovery cohort: 2162 participants and replication cohort: 724 women) identified multiple metabolomics that were highly correlated with chronological age (Yu et al., 2012). A cohort study of 2239 participants from the UK showed that metabolomic age developed based on metabolic profiles was associated with the prevalence of obesity, diabetes, heavy alcohol use and depression (Robinson et al., 2020). Several other studies have demonstrated that metabolomic age acceleration was associated with an increased risk of mortality (Deelen et al., 2019; Fischer et al., 2014), and CVD (van den Akker et al., 2020). As previous studies have focused on one or several diseases, whether metabolomic age acceleration is associated with the risk of a wide range of chronic diseases remains to be explored. These studies are also limited by the relatively small sample sizes or cross-sectional design, while biological clock estimates using machine learning can be improved by increasing the training sample size (Zhang et al., 2019).

Using the UK Biobank, we aimed to develop metabolomic age based on metabolomics measured by nuclear magnetic resonance spectroscopy using machine learning. We then examined the association between metabolomic age and a wide range of individual chronic diseases. The interplay between metabolomic age and age, sex, diet and metabolic disorders for chronic diseases was then examined.

# 2 | METHODS

### 2.1 | Study population

The present study was based on the UK Biobank, which is a population-based cohort of more than 500,000 participants aged 39–73 years at enrolment (Sudlow et al., 2015). Data on demographic factors, lifestyle and medical history were collected using self-administered questionnaires from 502,505 participants out of approximately 9.2 million invited people. Details of the study design have been shown elsewhere (Sudlow et al., 2015).

The UK Biobank Study's ethical approval has been granted by the National Information Governance Board for Health and Social Care and the NHS North West Multicenter Research Ethics Committee (REC reference: 16/NW/0274). All participants provided informed consent through electronic signature at recruitment.

### 2.2 | Ascertainment of diseases

Individual diseases were defined if participants reported that they had ever been told by a doctor that they had the disease (Field code: Table S1). Fifty major diseases (all important conditions of interest) such as cardiometabolic disorders (CMD) (including diabetes, coronary heart disease (CHD), heart failure, atrial fibrillation and stroke), cancer (including melanoma, lung cancer and stomach cancer), chronic obstructive pulmonary disease (COPD) and chronic kidney disease (CKD) were included in the analysis.

Inpatient data were used to identify additional disease cases at baseline. Inpatient hospital records were captured using the Hospital Episode Statistics database, the Scottish Morbidity Record, and the Patient Episode Database in England, Scotland and Wales (Sudlow et al., 2015). In the UK Biobank, the inpatient hospital data were available since 1997 (Sudlow et al., 2015). The international classification diseases codes for each of the 50 diseases are listed in Table S2. Incident cases of these diseases were defined using inpatient and mortality data. The onset date of diseases was defined as the earliest recorded date regardless of sources. Person-years for each disease were calculated from the date of baseline assessment to the date of onset, date of death or the end of follow-up (31 December 2020 for England and Wales and 31 January 2021 for Scotland), whichever came first.

# 2.3 | Metabolomic profiling

In the UK Biobank, metabolomic profiles were measured according to the structure and chemical properties of molecules using a highthroughput nuclear magnetic resonance-based metabolic biomarker profiling platform (Würtz et al., 2017). EDTA plasma samples were collected from a randomly selected subset of 117,121 participants at baseline and 5000+ participants at the first repeat visit. Venous blood sampling was collected and transported to a central laboratory and stored in ultra-low temperature archives. The measurements at baseline were used in our analysis. Measurements were conducted for 249 metabolic traits (168 concentrations and 81 ratios) including the lipoprotein lipids, fatty acids, amino acids, glycolysis, organic acids and nucleotides. Automated quality control was performed, and biomarker values substantially affected by interfering substances were removed (https://biobank.ctsu.ox.ac.uk/ crystal/label.cgi?id=220) (Soininen et al., 2015; Würtz et al., 2017). Metabolomic levels were normalized with mean = 0 and standard deviation (SD) = 1.

### 2.4 | Covariates

Demographic information on age, sex, ethnicity, education and annual household income was self-reported. Sleep duration was assessed based on the question 'About how many hours' sleep do you get in every 24 h?' Physical activity was assessed using a short form of the International Physical Activity Questionnaire. Diet score was computed based on seven commonly eaten food groups with a higher score representing a healthier diet (Lourida et al., 2019). Healthy diet was defined as diet score ≥4 and unhealthy diet as diet score <4 (Lourida et al., 2019). Medication use for antihypertension, lipid-lowering and glucose-lowering was self-reported. Body mass index (BMI) was calculated based on measured height and weight. A genetic risk score (GRS) for longevity was computed using 78 singlenucleotide polymorphisms with a higher score representing longer longevity (Timmers et al., 2020).

# 2.5 | Statistical analysis

We randomly selected 50% of the population with metabolomic data stratified by the assessment center to train (25% randomly



selected participants) and test (the remaining 25% participants) the chronological age prediction model based on 168 metabolomic profiling concentrations. Data from the remaining 50% participants were used to develop metabolomic age and examine the association between metabolomic age and the risk of multiple diseases. Multiple linear regression models with the chronological age as the dependent variable were used to develop metabolomic age. We selected Gaussian family distribution when establishing prediction model using machine learning (Figure 1). The hyper-parameters alpha and lambda specify the regularization strength and the regularization distribution between L1 (LASSO) and (ridge regression) L2 penalties, respectively. In this study, we used the R-square to determine the best prediction performance.

Age gap was calculated by subtracting chronological age from metabolomic age. Given the age gap for individuals with different ages might represent different metabolomic ageing levels, chronological age-adjusted age gap (CAAG) was calculated with the use of regression models (Willett et al., 1997). Baseline characteristics were expressed as frequency (%) or means  $\pm$  SDs. ANOVA for continuous variables and chi-square test for categorical variables were used to test the difference of characteristics by quintiles of CAAG.

The association between CAAG and incidence of each chronic disease was examined using the Cox proportional regression models. For each individual chronic disease, participants with the corresponding disease at baseline were excluded from the analysis. Three models were tested: (1) Model 1 was unadjusted; (2) Model 2 was adjusted for age, sex, ethnicity, education, household income, diet score, alcohol consumption, physical activity, smoking, sleep duration, fasting duration and GRS for longevity; (3) Model 3 was adjusted for Model 2 plus BMI, high cholesterol, hypertension, diabetes and antihypertensive, glucose-lowering, and lipid-lowering medications. CAAG was analysed in quintiles as well as a continuous variable (each year). Benjamin-Hochberg's procedure was used to control the false discovery rate at a 5% level for multiple comparisons (Benjamini & Hochberg, 1995).

Sensitivity analysis of the association between CAAG and risk of individual diseases was conducted among individuals by excluding those developed the disease in the first year of follow-up or by excluding those developed in the first 5 years of follow-up. Whether associations between CAAG and chronic diseases were modified by age, sex, education, diet quality, metabolic disorders or GRS for longevity was tested using the Cox proportional regression models.

Percentages of individuals with missing data in BMI, physical activity, income and education were 2%, 19%, 14% and 1%, respectively. Given that individuals with missing data in outcome/exposure variables were excluded from the analysis, multiple imputations for missing data in covariates only using the fully conditional specification method were conducted to create 10 imputed datasets.

Data analyses were conducted using SAS 9.4 for Windows (SAS Institute Inc.) and all p-values were two-sided with statistical significance set at <0.05.



FIGURE 1 Flowchart for the development of metabolomic age and its association with the risk of chronic diseases (a) refers to the division of datasets; (b) refers to the development of metabolomic age and its linkage to chronic diseases.

# 3 | RESULTS

### 3.1 | Population selection

Individuals with no data on metabolomic profiles (n=391,785), or with missing data on inpatient data (n=18) were excluded from the analysis. We included 110,692 participants in the final analysis. The analysis for the association between metabolic age and chronic diseases was conducted in the validation dataset (n=55,346,54.1% female, aged 39–71 [mean±SD:  $56.5\pm8.1$ ] years at baseline).

### 3.2 | Metabolomic age and chronological age

In the machine learning analysis, the LASSO and ridge regression model with  $\alpha = 0.2$  and  $\lambda = 0.0001$  showed the best prediction performance in the training and testing datasets. The algorism was then used to predict metabolomic age in the validation dataset (Figure 1).

The distribution of predicted metabolomic age is shown in Figure 2. The correlation between metabolomic age and chronological age was 0.56 (95% CI: 0.55–0.56). The mean gap between metabolomic age and chronological age was 0.02 (SD=6.93) years and the mean absolute error was 5.52 (SD=4.27) years.

The age gap decreased with increasing age and similar trends were seen in both women and men. The age gap was  $5.1\pm5.1$  years in individuals aged younger than 60 years and  $-5.0\pm4.4$  years in those aged 60 years or older (Figure S1). Therefore, CAAG was used in the association analysis.

Top 10 leading determinants of metabolomic age included omega-3 fatty acids, docosahexaenoic acid, citrate, triglycerides in

intermediate-density lipoprotein, triglycerides in large low-density lipoprotein, triglycerides in very small very low-density lipoprotein, triglycerides in large high-density lipoprotein, triglycerides in lowdensity lipoprotein, tyrosine and saturated fatty acids (Figure S2). Similar metabolomic profiles that were the strongest predictors of chronological age were found (Figure S3).

### 3.3 | Baseline characteristics

Individuals with greater CAAG were more likely to be female and have lower household income and higher BMI and a higher prevalence of hypertension (Table 1).

# 3.4 | Chronological age-adjusted age gap and incidence of individual diseases

Given the difference in the number of cases at baseline between individual diseases, the follow-duration differed between these diseases. The median follow-up duration ranged from 11.2 years for dyspepsia to 11.9 years for multiple sclerosis. The number of incident cases ranged from 46 for multiple sclerosis to 6484 for dyspepsia.

After controlling for false discovery rate, CAAG was significantly associated with the incidence of 25 out of 50 individual chronic diseases in Model 1. After adjustment for demographic factors, lifestyle, fasting duration, metabolic disorders and mediations use for hypertension and high cholesterol, the association with 15 chronic diseases remained significant. Each year increment of CAAG was associated with a 1% (95% Cl: 0%–2%), 3% (1%–4%), 2% (1%–4%), 3% (1%–5%),



**FIGURE 2** Distribution of metabolomic age in women and men. Age gap was calculated by subtracting chronological age from metabolomic age. Chronological age-adjusted age gap was calculated with use of regression models. Red curves represent the trend of the distribution.

2% (0.3%–4%) and 2% (1%–3%) higher risk of CHD, heart failure, atrial fibrillation, stroke and other cardiac disease, respectively. Greater GAAG was associated with an increased risk of diabetes (HR (95% Cl): 1.05 [1.03–1.06]), hypertension (1.01 [1.00–1.02]), dyslipidemia (1.02 [1.01–1.03]), oesophageal cancer (1.06 [1.02–1.11]), other cancers (1.01 [1.00–1.02]), alcohol use disorder (1.04 [1.01–1.06]), chronic liver disease (1.07 [1.03–1.11]), COPD (1.02 [1.01–1.03]), CKD (1.05 [1.04–1.06]) and age related macular degeneration (1.02 [1.00–1.04]).

The association between CAAG and risk of anxiety, asthma, diverticulitis, thyroid disorders and eczema was attenuated to be nonsignificant after adjustment for metabolic disorders and the use of related medications (Figure 3).

This was inconsistent with the results when CAAG was analysed in quintiles as categorical variables (Table S3). The survival plots for multiple individual diseases by quintiles of CAAG with significant associations are shown in Figure S4.

# 3.5 | Moderation analysis

The association between CAAG and risk of peripheral vascular disease, other cardiac diseases, fracture, cataract and thyroid disorder was stronger among individuals with unhealthy diet. Greater CAAG was associated with a lower risk of prostate disorders in individuals with healthy diet (HR [95% CI] for each year increment: 0.95 [0.92-0.99]) but not those with unhealthy diet (1.00 [0.99-1.01], Figure 4).

The association between CAAG and risk of heart failure, other cardiac diseases, diabetes, hypertension, dyslipidemia, COPD and CKD was stronger among younger than in older individuals. Greater CAAG was associated with increased risk of depression, schizophrenia, dementia, osteoporosis, lung cancer and cataract in younger individuals only (Figure 5).

The association between CAAG and risk of some CVDs, some neurological disorders, and age related macular degeneration was stronger in individuals with lower education (Figure 6).

The association between CAAG and risk of melanoma and chronic liver disease was stronger among individuals without hypertension than those with hypertension. Larger CAAG was associated with an increased risk of peripheral vascular disease, inflammatory bowel disease and cataract was significant among those with diabetes only. The association between CAAG and risk of epilepsy, hypertension and dyslipidemia was stronger in individuals with low GRS for longevity than those with high GRS (Figure S5).

# 3.6 | Sensitivity analysis

Similar results for the association between chronological ageadjusted age gap and risk of individual diseases were seen among Age (years)

Ethnicity White Non-White Education High Intermediate Low

Household income (pounds)

<18,000 18,000-30,999 31,000-51,999 52.000-100.000 >100,000 Alcohol consumption

> Never Previous Current

Smoking

Never Former Current

Diet score<sup>c</sup> Unhealthy diet Healthy diet Sleep duration (hours)

BMI (kg/m<sup>2</sup>) Genetic risk score for longevity<sup>d</sup>

Hypertension High cholesterol Diabetes

Physical activity

(MET-minutes/week)

Fasting duration (hours)

Antihypertensive medication Lipid lowering medication

Glucose lowering medication

Sex Women Men

TABLE 1 Baseline character

ina Cell	( <b>X</b> )				SHANG ET AL
stics across quinti	les of chronological	age-adjusted age	gap.		
Chronological ag	e-adjusted age gap (y	vears) <sup>a</sup>			
Quintile 1	Quintile 2	Quintile 3	Quintile 3 Quintile 4 Quintile 5		p-value <sup>b</sup>
(≤2.90)	(-2.90, -0.92)	(-0.93, 0.80)	(0.81, 2.9)	(>2.9)	—
56.6±8.2	56.5±8.2	$56.5 \pm 8.1$	$56.5 \pm 8.1$	56.6±8.0	1.00
4994 (45) 6076 (55)	5645 (51) 5423 (49)	5959 (54) 5110 (46)	6373 (58) 4696 (42)	6954 (63) 4116 (37)	<0.0001
10,320 (93) 750 (7)	10,410 (94) 658 (6)	10,519 (95) 550 (5)	10,526 (95) 543 (5)	10,452 (94) 618 (6)	<0.0001
3916 (35) 5401 (49) 1753 (16)	3787 (34) 5452 (49) 1829 (17)	3671 (33) 5623 (51) 1775 (16)	3692 (33) 5556 (50) 1821 (17)	3589 (32) 5581 (51) 1900 (17)	0.00011
2329 (21) 3189 (29) 2932 (27) 2088 (19) 532 (5)	2326 (21) 3164 (29) 2933 (26) 2124 (19) 521 (5)	2414 (22) 3158 (28) 2855 (26) 2094 (19) 548 (5)	2449 (22) 3197 (29) 2842 (26) 2067 (19) 514 (5)	2798 (25) 3329 (30) 2667 (24) 1851 (17) 425 (4)	<0.0001
482 (5) 374 (3) 10,214 (92)	477 (4) 418 (4) 10,173 (92)	458 (4) 366 (3) 10,245 (93)	474 (4) 349 (3) 10,246 (93)	529 (5) 460 (4) 10,081 (91)	0.0323
6115 (55) 3777 (34) 1178 (11)	6058 (55) 3852 (35) 1158 (10)	6158 (56) 3806 (34) 1105 (10)	6007 (54) 3908 (35) 1154 (11)	5926 (53) 3973 (36) 1171 (11)	0.0277
2743±2808	2679±2734	$2623 \pm 2626$	$2588 \pm 2611$	$2512 \pm 2653$	<0.0001
4465 (40) 6605 (60)	4341 (39) 6727 (61)	4202 (38) 6867 (62)	4035 (36) 7034 (64)	3903 (35) 7167 (65)	<0.0001
$7.13 \pm 1.07$	$7.14 \pm 1.08$	$7.15 \pm 1.10$	$7.18 \pm 1.12$	$7.19 \pm 1.20$	< 0.0001
$26.47 \pm 4.06$	27.00±4.27	$27.45 \pm 4.50$	$27.79 \pm 4.91$	$28.64 \pm 5.52$	<0.0001
$0.49 \pm 0.05$	0.0195				
$3.80 \pm 2.50$	$3.79 \pm 2.47$	$3.74 \pm 2.31$	$3.75 \pm 2.40$	$3.84 \pm 2.45$	0.48
2434 (22)	2607 (24)	2863 (26)	3158 (29)	3738 (34)	<0.0001
1805 (16)	2022 (18)	2361 (21)	2559 (23)	3206 (29)	<0.0001
322 (3)	378 (3)	445 (4)	599 (5)	1130 (10)	<0.0001
1809 (16)	1969 (18)	2168 (20)	2486 (23)	3145 (28)	<0.0001
1286 (12)	1546 (14)	1807 (16)	2142 (19)	2962 (27)	< 0.0001
210 (2)	269 (2)	315 (3)	456 (4)	873 (8)	<0.0001

Note: Data are means  $\pm$  standard deviations, or N (%).

Abbreviations: BMI, body mass index; MET, metabolic equivalent.

<sup>a</sup>Age gap was calculated by subtracting chronological age from metabolomic age. Chronological age-adjusted age gap was calculated with use of regression models.

<sup>b</sup>ANOVA for continuous variables and chi-squared for categorical variables were used to test the difference of baseline characteristics across quintiles of chronological age-adjusted age gap.

<sup>c</sup>Diet score was computed based on seven commonly eaten food groups following recommendations on dietary priorities for cardiometabolic health with a higher score representing a healthier diet.

<sup>d</sup>Genetic risk score was calculated for longevity was calculated using 78 single-nucleotide polymorphisms.

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Disease	Hazard ratio (95% Cl), Model 1 <sup>*</sup>		Hazard ratio (9	5% CI), Model 2	Hazard ratio (95% CI), Model 3	
Cardiovascular disease						
Coronary heart disease	-	1.02 (1.01-1.03)	-	1.02 (1.01-1.03)	-	1.01 (1.00-1.02
Heart failure		1.04 (1.03-1.06)		1.03 (1.02-1.05)		1.03 (1.01-1.04
Atrial fibrillation		1.02 (1.01-1.04)		1.02 (1.01-1.04)		1.02 (1.01-1.04
Other cardiac problem	-	1.03 (1.02-1.04)	-	1.03 (1.02-1.04)	-	1.02 (1.01-1.03
Stroke		1.03 (1.01-1.05)		1.03 (1.01-1.05)		1.03 (1.01-1.05
Peripheral vascular disease	- <b>-</b> -	1.02 (1.00-1.04)		1.03 (1.00-1.05)		1.01 (0.99-1.04
Metabolic disorder						
Hypertension	-	1.03 (1.02-1.03)	-	1.02 (1.01-1.03)	-	1.01 (1.00-1.02
Diabetes		- 1.08 (1.07-1.09)	-	1.06 (1.05-1.07)		1.05 (1.03-1.06
Dyslipidemia	-	1.05 (1.04-1.06)	-	1.05 (1.04-1.05)	-	1.02 (1.01-1.03
Cancer						
Non-melanoma skin cancer	<b>.</b>	1.00 (0.99-1.01)		1.01 (0.99-1.02)	-	1.01 (1.00-1.02
Melanoma	<b>_</b>	1.00 (0.97-1.03)	<b></b>	1.00 (0.97-1.03)		1.01 (0.97-1.04
Lung cancer		1.01 (0.98-1.03)		1.01 (0.99-1.04)		1 01 (0 99-1 04
Stomach cancer -	_ <b></b>	0.94 (0.90-0.99)		0.95 (0.91-1.00)		0.96 (0.91-1.01
Oesophageal cancer		- 1.04 (1.00-1.09)		1.05 (1.01-1.10)		1.06 (1.02-1.11
Colon cancer		0.09 (0.06 1.01)		0.09 (0.96 1.01)		0.99 (0.96 1.01
Rootal cancer		1.01 (0.98 1.05)		1.03 (0.90-1.01)		1.02 (0.98 1.06
		1.01 (0.98-1.03)		1.03 (0.99-1.07)		1.02 (0.98-1.00
Prostate cancer		0.98 (0.97-1.00)		0.98 (0.96-1.00)		0.99 (0.97-1.00
Ovarian cancer*		0.98 (0.94-1.03)		0.98 (0.94-1.03)		0.98 (0.94-1.03
Breast cancer*		1.01 (0.99-1.03)		1.01 (0.99-1.02)		1.01 (0.99-1.02
Other cancers	<b>-</b>	1.01 (1.00-1.02)		1.01 (1.00-1.02)		1.01 (1.00-1.02
Neurological disorder						
Dementia	- <b>-</b>	1.02 (1.00-1.04)		1.02 (1.00-1.04)		1.01 (0.99-1.03
Parkinson's disease		0.99 (0.96-1.02)		0.99 (0.96-1.03)		1.00 (0.97-1.03
Multiple sclerosis		- 1.02 (0.94-1.10) -		- 0.99 (0.92-1.08) -		- 1.00 (0.92-1.08
Migraine	÷■−	1.02 (0.99-1.04)	<b></b>	1.00 (0.97-1.03)		0.99 (0.96-1.02
Epilepsy	<b>⊹</b> ∎	1.02 (0.99-1.06)		1.03 (0.99-1.06)		1.02 (0.99-1.05
Psychological disorder						
Depression		1.03 (1.01-1.04)		1.01 (0.99-1.02)		1.00 (0.98-1.02
Anxiety		1.03 (1.01-1.04)	<b>—</b> -	1.01 (1.00-1.02)	- <b>+</b>	1.00 (0.99-1.02
Alcohol use disorder		1.04 (1.01-1.06)		1.04 (1.02-1.07)		1.04 (1.01-1.06
Psychoactive substance abuse	<b></b>	1.03 (0.98-1.07)	·	- 1.04 (1.00-1.08)	÷	- 1.04 (0.99-1.08
Schizophrenia		- 1.06 (1.01-1.11)	÷=	- 1.03 (0.99-1.09)	<b>_</b>	- 1.02 (0.97-1.08
Digestive disorder						
Dyspepsia	-	1.01 (1.00-1.02)	÷	1.00 (0.99-1.01)	- <b>-</b>	1.00 (0.99-1.01
Treated constipation	-	1.01 (1.00-1.02)	<b></b>	1.00 (0.99-1.02)	- <b>-</b>	1.00 (0.99-1.01
Diverticular disease	-	1.02 (1.01-1.03)	-	1.01 (1.00-1.02)	-	1.01 (1.00-1.02
Inflammatory bowel disease	<b></b>	1.00 (0.97-1.03)	<b>_</b>	1.00 (0.97-1.04)	<b>_</b>	1.00 (0.97-1.04
Irritable bowel syndrome		1.02 (1.00-1.04)	<b></b>	1.00 (0.98-1.02)	<b></b>	1.00 (0.98-1.02
Chronic liver disease				1.08 (1.04-1.12)		1.07 (1.03-1.11
Respiratory disorder		,		)		
COPD	-	1.02 (1.01-1.04)	-	1.02 (1.01-1.03)	-	1.02 (1.01-1.03
Asthma		1.02(1.01-1.05)		1.02(1.01-1.03)	_	1.01 (1.00-1.03
Bronchiactasis		1.00 (0.98 1.03)		1.01 (0.98, 1.03)		1.01 (0.98.1.02
Chronic hidrox diagon		1.00 (0.98-1.03)		1.01 (0.98-1.03)		1.01 (0.98-1.04
Museuleslestel disease	-	1.07 (1.00-1.08)		1.00 (1.03-1.07)		1.03 (1.04-1.00
Musculoskeletal disorders	_	1.02 (1.01.1.02)	_	0.00 (0.00 1.00)	<u>i</u>	1 00 (0 00 1 0)
Osteoporosis	<b>_</b>	1.02 (1.01-1.03)	<u> </u>	0.99 (0.99-1.00)	-	1.00 (0.99-1.0)
Fracture		1.00 (0.98-1.03)		1.00 (0.97-1.02)		1.00 (0.98-1.02
Ophthalmic condition	- -		-		1	
Glaucoma		1.00 (0.98-1.02)		1.00 (0.98-1.02)		1.00 (0.98-1.02
Cataract		1.01 (1.00-1.01)	-	1.00 (0.99-1.00)		1.00 (0.99-1.00
AMD		1.03 (1.01-1.05)		1.02 (1.00-1.04)		1.02 (1.00-1.04
Other disease						
Thyroid disorders	-	1.04 (1.03-1.06)	<b>⊢</b> ∎-	1.02 (1.00-1.03)	-	1.01 (1.00-1.0
Psoriasis/eczema		1.03 (1.01-1.05)	<b>⊢∎</b>	1.02 (1.00-1.04)		1.01 (0.99-1.04
Meniere's disease		0.98 (0.92-1.05)		0.96 (0.92-1.01)	-	0.96 (0.90-1.03
Prostate disorders <sup>†</sup>	÷	1.00 (0.99-1.01)	-	0.99 (0.98-1.00)	-	0.99 (0.98-1.00
	_'	0.00 (0.01.1.04)	-	0.00 (0.02, 1.02)	-	0.04 (0.01.1.01

- -

FIGURE 3 Legend on next page

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FIGURE 3 The association between each year increment in chronological age-adjusted age gap and risk of individual diseases in the validation population. Age gap was calculated by subtracting chronological age from metabolomic age. Chronological age-adjusted age gap was calculated with the use of regression models. \*Cox proportional regression models were used to examine the association between chronological age-adjusted age gap (each year increment) and incidence of individual chronic diseases. Model 1 was unadjusted; Model 2 was adjusted for Model 1 plus age, sex, ethnicity, education, household income, diet score, alcohol consumption, physical activity, smoking, sleep duration, fasting duration, and GRS for longevity; Model 3 was adjusted for Model 2 plus BMI, high cholesterol, hypertension, and antihypertensive and lipid-lowering medications (hypertension or antihypertensive medication use at baseline was not adjusted for the analysis of incident hypertension given these participants with hypertension or antihypertensive medication use were excluded from the analysis). Red color squares refer to significantly positive associations while green color squares refer to significantly inverse associations. The significant associations in Model 1 were defined as *p*-value<0.05 after adjustment for false discovery rate. <sup>†</sup>These analyses were conducted among women only. AMD, age related macular degeneration; CI, confidence interval; COPD, chronic obstructive pulmonary disease; HR, hazard ratio.

Diama	Hazard ratio (95% CI) for each 1-year increment of		
Disease	chronological age-adjusted age gap	interaction	
Peripheral vascular disease		0.0135	
Unhealthy diet	<b>———</b> 1.07 (1.04-1.10)		
Healthy diet	1.02 (0.99-1.05)		
Other cardiac disease		0.0084	
Unhealthy diet	<b>———</b> 1.06 (1.04-1.07)		
Healthy diet	<b>—</b> 1.03 (1.02-1.05)		
Fracture		0.0010	
Unhealthy diet	<b>—</b> 1.05 (1.01-1.09)		
Healthy diet	0.98 (0.95-1.01)		
Cataract		0.0142	
Unhealthy diet	1.02 (1.00-1.03)		
Healthy diet	1.00 (0.99-1.01)		
Thyroid disorder		0.0011	
Unhealthy diet	<b>———</b> 1.06 (1.03-1.08)		
Healthy diet	1.01 (0.99-1.03)		
Prostate disorders		0.0108	
Unhealthy diet	<b>-+</b> - 1.00 (0.99-1.01)		
Healthy diet	0.95 (0.92-0.99)		
0.92	0.96 1.00 1.04 1.08		

FIGURE 4 The association between chronological age-adjusted age gap and incidence of chronic diseases moderated by diet score. Cox proportional regression models were used to test whether diet quality modified the association between chronological age-adjusted age gap and incidence of chronic diseases. Only the results with significant interaction are shown in this figure. Horizontal lines indicate the range of the 95% confidence interval. The vertical dash lines represent the hazard ratio of one.

individuals by excluding those developed the disease in the first year of follow-up (Figure S6, Table S4) or by excluding those developed in the first 5 years of follow-up (Figure S7, Table S5).

# 4 | DISCUSSION

Using data from this large cohort study, we found greater CAAG was associated with an increased risk of 24 individual chronic diseases. The association with 15 diseases including most CVDs, all metabolic disorders, some cancers, alcohol use disorder, CKD, respiratory disorders, chronic liver disease and age related macular degeneration remained significant after adjustment for metabolic disorders and the use of related medications at baseline. The association with some CVDs, fracture, thyroid disorders and cataract was stronger among individuals with unhealthy diet. Greater CAAG was associated with lower risk of prostate disorders among individuals with healthy diet only. The association between CAAG

and some chronic diseases was stronger among younger individuals, lowly educated individuals or those with metabolic disorders/ low GRS of longevity.

The strong associations between metabolism and ageing provide the rationale for examining metabolomic clocks (López-Otín et al., 2016). Metabolomic profiles including fatty acids, lipids and amino acids have been demonstrated to be strong predictors of longevity (Gonzalez-Covarrubias et al., 2013). Consistently, we found fatty acids, amino acids and triglycerides (in intermediate-density lipoprotein, large low-density lipoprotein, very small very lowdensity lipoprotein, large high-density lipoprotein and low-density lipoprotein) were among leading determinants of metabolomic age. Recent evidence suggests machine learning based on a larger training sample size has improved the precision of epigenetic clock estimates (Zhang et al., 2019). Our study showed that machine learning based LASSO and ridge regression analysis provided fair prediction of chronological age based on metabolomics. Noise in metabolomic data may limit the utility of metabolomic age (Rutledge et al., 2022),

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FIGURE 5 The association between chronological age-adjusted age gap and incidence of chronic diseases moderated by chronological age. Cox proportional regression models were used to test whether chronological age modified the association between chronological age-adjusted age gap and incidence of chronic diseases. Only the results with significant interaction are shown in this figure. Horizontal lines indicate the range of the 95% confidence interval. The vertical dash lines represent the hazard ratio of one. CKD, chronic kidney disease; CI. confidence interval: COPD, chronic obstructive pulmonary disease.



such that previous studies on metabolomic age only focused on several conditions including mortality, CVDs, obesity, diabetes and psychological disorders (Deelen et al., 2019; Fischer et al., 2014; Hertel et al., 2016; Menni et al., 2013; Robinson et al., 2020). Whether metabolomic age is associated with the risk of a wide range of chronic diseases needs to be explored in more cohort studies.

The important role of metabolomic profiles on CMDs has been highlighted in previous studies (Buergel et al., 2022; Shah

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Disease	Hazard ratio (95% CI) for each 1-year increment of chronological age-adjusted age gap				
Atrial fibrillation		0.0015			
High education	1.01 (0.98-1.03)				
Intermediate education	<b>——</b> 1.06 (1.04-1.08)				
Low education	<b>—</b> 1.06 (1.04-1.09)				
Stroke		0.0228			
High education	0.99 (0.95-1.03)				
Intermediate education	<b>———</b> 1.05 (1.02-1.08)				
Low education	<b>—</b> 1.06 (1.02-1.10)				
Other cardiac disease		0.0096			
High education	1.02 (1.00-1.04)				
Intermediate education	<b>——</b> 1.05 (1.03-1.06)				
Low education	1.05 (1.03-1.07)				
Epilepsy		0.0063			
High education	0.97 (0.92-1.03)				
Intermediate education	1.03 (0.99-1.08)				
Low education	1.07 (1.01-1.13)				
Parkinson's disease		0.0158			
High education	0.94 (0.88-1.00)				
Intermediate education	1.01 (0.96-1.06)				
Low education	1.04 (0.98-1.10)				
AMD		0.0355			
High education	0.99 (0.95-1.02)				
Intermediate education	1.03 (1.01-1.07)				
Low education	1.03 (0.99-1.07)				
0.8	8 0.92 0.96 1.00 1.04 1.08 1.12				

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FIGURE 6 The association between chronological age-adjusted age gap and incidence of chronic diseases moderated by education. Cox proportional regression models were used to test whether education modified the association between chronological age-adjusted age gap and incidence of chronic diseases. Only the results with significant interaction are shown in this figure. Horizontal lines indicate the range of the 95% confidence interval. The vertical dash lines represent the hazard ratio of one. AMD, age-related macular degeneration; CI, confidence interval.

et al., 2012). We found accelerated CAAG was associated with an increased risk of most CMDs of interest including CHD, heart failure, stroke and diabetes. Although these associations were independent of metabolic disorders and the use of medications for hypertension and lipids at baseline, the adjustment for these covariates attenuated the associations. This suggests good management of metabolic disorders may help mitigate the adverse effect size of accelerated metabolomic age on CMDs. It is also possible that the intake of respective medications might reduce the validity of the metabolomics thus biasing the algorithm of metabolomic age. In a cross-sectional analysis of data from 26 community and hospital-based cohorts, larger age gap (metabolomic age minus chronological age) was associated with a higher likelihood of diabetes (OR [95% CI] for each 10-year increment of age gap: 2.52 [1.93-3.11]) among 12,633 participants (van den Akker et al., 2020). In the same study, a longitudinal analysis of 5410 participants with a mean follow-up of 3.3 years showed that larger age gap was associated with an increased risk of CHD (HR [95% CI] for each 10-year increment: 1.25 [1.11-1.40]) (van den Akker et al., 2020). Data from the UK Airwave cohort demonstrated that metabolomic age was corrected with obesity and diabetes (Robinson et al., 2020). Our findings regarding heart failure, atrial fibrillation, stroke, peripheral vascular disease and other cardiac disease need to be confirmed by future cohort studies with long follow-up duration.

As metabolomic state has been linked to multiple diseases (Buergel et al., 2022), it is also of great interest to examine whether metabolomic age is predictive of many other diseases rather than CMDs only. Robinson et al. reported that metabolomic age acceleration was associated with heavy alcohol use and depression (Robinson et al., 2020). Likely, we found greater CAAG was associated with higher risk of incident depression, anxiety, alcohol use disorder and psychoactive substance abuse. The association for depression and anxiety was attenuated to be non-significant after adjustment for metabolic disorders and antihypertensive and lipid-lowering medications indicating that good control of traditional metabolic biomarkers might contribute to the reduction of

the risk due to metabolomic age acceleration. A cross-sectional analysis of data from the Study of Health in Pomerania showed that metabolic age acceleration was associated with kidney malfunction (microalbuminuria, albumin-to-creatinine ratio) (Hertel et al., 2016). A recent longitudinal multi-omics study in humans revealed that kidney dysfunction was involved in the process of ageing (Ahadi et al., 2020). This is consistent with our study demonstrating that CAAG was associated with an increased risk of CKD. In a cohort study of 6055 individuals from the UK, an age-related metabolite C-glycosyl tryptophan was associated with lung function (forced expiratory volume) (Menni et al., 2013). We found greater CAAG was associated with an increased risk of COPD and asthma and the association for COPD was even independent of metabolic disorders and related medications use. The longitudinal multi-omics study also identified an ageing pathway related to liver dysfunction (Ahadi et al., 2020). Consistently, metabolomic age was strongly associated with the risk of liver disease in our study. We also found larger CAAG was associated with a higher risk of oesophageal cancer but not other cancers. A recent systematic review reported that a number of metabolites were identified for oesophageal cancer but the results were inconsistent between previous studies (Huang et al., 2020). Meanwhile, metabolomic age acceleration was associated with an increased risk of dyspepsia, diverticulitis, osteoporosis and thyroid disorders before but not after adjustment for metabolic biomarkers and use of antihypertensive and lipid-lowering medications in our study. Our study developed a metabolomic age that was independently predictive of multiple chronic diseases including psychological disorders, CKD, COPD, chronic liver disease and oesophageal cancer, which might be useful for the screening and prevention of these diseases.

The importance of diet in the development of chronic diseases should not be overlooked (Schulze et al., 2018; Shan et al., 2020). In moderation analysis, the association between CAAG and the risk of several types of CVD, CKD and chronic liver disease was weaker among individuals with healthy diet. Therefore, individuals are recommended to accommodate healthy diet habits to mitigate the risk of metabolomic ageing. Stronger associations between metabolomic age and some chronic diseases seen in individuals with metabolic disorders in our study highlight the importance of the management of metabolomics in these people. The association between CAAG and the risk of some chronic diseases was more pronounced among young than older individuals. This is consistent with previous studies showing that metabolic disorders diagnosed at younger age was associated with greater risk of CVD, dementia and mortality (Shang et al., 2021, 2022; Zhao et al., 2021). The association between CAAG and the risk of some chronic diseases was stronger among individuals with lower education. This is possibly due to the fact that individuals with higher education are more likely to seek health care and less likely to develop chronic diseases with metabolomic ageing (Brayne et al., 2010; Livingston et al., 2020). We also found that greater CAAG was associated with an increased risk of epilepsy, hypertension and dyslipidemia in individuals with high GRS of longevity. One

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possible explanation for this is that low GRS of longevity was associated with higher prevalence of metabolic disorders and higher genetic risks of epilepsy, hypertension and dyslipidemia were detected in long-lived individuals (Hu et al., 2022). Metabolomic age provides different prediction values of some chronic diseases between diet, education, metabolic disorders, age or GRS groups.

This is the first study with a large sample size and long follow-up duration to develop metabolomic age and examine its association with a wide range of chronic diseases. Several potential limitations need to be considered in our study. First, the metabolomic age was developed based on metabolomics measured at one time point, such that metabonomic dynamics with ageing within an individual could not be estimated. Second, the plasma sample in the UK Biobank was non-fasting, which might bias the associations. However, the adjustment for fasting duration in the analysis did not substantially change the association between CAAG and chronic diseases. Third, most of the participants in our analyses were Caucasians thus our findings may not be generalized to other ethnic groups. Finally, the number of incident cases for several chronic diseases (such as multiple sclerosis) was small, which might have reduced the statistical power to test significance.

In conclusion, metabolomic age plays an important role in the development of multiple chronic diseases including CMDs, psychological disorders, COPD, CKD, liver disease and some cancers. Healthy diet may help mitigate the risk for some chronic diseases due to metabolomic age acceleration. Age, education, metabolic disorders and GRS for longevity may modify the association between metabolomic age and some chronic diseases. Our findings may help facilitate the understanding of ageing process related to metabolomics thus enhancing healthy ageing.

### AUTHOR CONTRIBUTIONS

XS, MH, and HY conceived and designed the study. ZZ, WW performed data curation. XS conducted data analysis and drafted the initial manuscript. XS, JL, XLZ, YH, SL, ZZ, WW, XYZ, ST, YH, ZG, HY, and MH made a critical revision to the manuscript for important intellectual content. All authors read the manuscript and approved the final draft.

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### CONFLICT OF INTEREST STATEMENT

The authors declare that they have no competing interests.

### DATA AVAILABILITY STATEMENT

Data are available in a public, open access repository (https://www. ukbiobank.ac.uk/).

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### SUPPORTING INFORMATION

Additional supporting information can be found online in the Supporting Information section at the end of this article.

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