

**Downtime of in-use water pump installations for high-rise residential
buildings**

Short title: Downtime of in-use water pumps

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Abstract

This study examines the fresh and flushing water pump installations for high-rise residential buildings in Hong Kong in terms of system availability, mean time to failure, mean time between failures and restoration time. Together with some reliability data published elsewhere, it applies Bayesian analysis to improve our understanding of the downtime characteristics of water pump installations. For three consecutive years (2005-2007), water pump failures in 46 typical high-rise residential buildings were recorded to determine the component failure rates. In order to study the failure patterns, Monte Carol simulations were performed for the operations of 100 parallel pump sets over a period of 10 years. The mean time to failure, total downtime, failure counts and system availability estimated for the fresh water pump installations were 1.24 years, 8990 hours, 709 and 90% while those estimated for the flushing seawater pump installations were 0.46 years, 4049 hours, 2081 and 78%, respectively. The results are useful in the calculation of water supply availability for high-rise residential buildings while keeping the balance between maintenance cost and system reliability. This study also demonstrates a method for reliability modelling of water supply for high-rise residential buildings.

Keywords

Water pump installation; reliability; seawater flushing; water supply

Practical applications

This study demonstrates a method for reliability modelling of building services systems. With the use of Bayesian analysis, example estimates of the mean time to failure, total downtime, failure counts and system availability were determined for the fresh and flushing water pump installations for high-rise residential buildings in Hong Kong.

List of symbols

A_i	a set of mutually exclusive and exhaustive events A describing the existing understanding of failure rates i , for $i=1,2,\dots$
B	an event of new failure-free observation
$G()$	lognormal distribution function
G_m	Gamma distribution with explanatory parameters β_1 and β_2
GM, GSD	geometric mean and geometric standard deviation
M	total number of installation components of the same type
N_c	number of components connected in series
N_f	total failure counts in a period
N_o	total number of time periods of faulty-available pairs for an installation
p	statistic p-value of a specified test
$P()$	probability
S, S_i	component operation state and operation state of component i
SD	standard deviation
x	percentile value

Greek

ϑ	a random number between 0 and 1
α_0, α_1	installation availability and unavailability
β_1, β_2	explanatory parameters for Gamma distribution function
λ	failure rate (h^{-1})
λ_a, λ_a'	prior and posterior failure rates (h^{-1})
μ	mean
ν, ν_0	regression constants
σ	standard deviation
τ	service period (h)
τ_0, τ_1	time between failures (h) and downtime (h)
τ_{2L}	downtime of 2 pump installations in parallel (h)
τ_m	mean time to failure (h)
τ_∞	operation period of a pump installation (h)
ζ	a normally distributed parameter with mean μ and standard deviation σ

Subscripts

a, a'	of prior value and posterior value
i	of i-th value
L, 2L	of a pump installation, of 2 pump installations in parallel
max	of maximum value
min	of minimum value
pro	of probable value

Superscripts

~	of a distribution
*	of a value

1. Introduction

Hong Kong is a high-rise city in which water is supplied through two completely separate networks – one for fresh water supply and the other for seawater flushing. The latter is a cost-effective and environmentally friendly water conservation measure.¹ As water mains below the street operate at different pressures, pumping facilities are usually provided at each point where the water enters a high-rise residential building. These pumping

facilities are installed in duplicate to permit continuous water supply when one of the pumps fails. Water secured from the street mains is stored in a break tank and then transferred through a pair of transfer pumps to a gravity tank elevated above the building roof for water distribution to every floor of the building. Sufficient water pressure is ensured by a pair of booster pumps set up especially for the topmost floors. Sometimes, instead of a gravity tank system, a hydro pneumatic pressure boosting system or a variable volume pumping system is used.

The usual problems associated with city mains are deteriorated pipes, defective joints, corrosion and faulty service connections. Among them, corrosion appears to be a detrimental factor in pipe failures.² In Hong Kong, water installation equipment and materials must be up to standard. However, the current standards are according to the test reports issued by the manufacturers. Systematic considerations for the influence of various installation arrangements and usage patterns on the reliability of a water supply system in buildings are still lacking. A study showed that a geographic information system (GIS) model with historical repair data, soil type and temperature can be used to identify the areas of a water distribution network where failure risk exists.³

Together with some reliability data published elsewhere, this study applied Bayesian analysis to improve our understanding of the downtime characteristics of water pump installations.

2. Water supply systems

A survey on water pump failures in high-rise residential buildings was conducted in Hong Kong from 2005 to 2007. It covered the three major categories of local high-rise housing estates: (1) publicly owned; (2) government subsidized; and (3) privately owned.⁴ A total of 46 buildings, whose details are outlined in Table 1, were surveyed.

Figure 1 illustrates a typical water pump installation for the surveyed buildings. The installation consists of two gate valves, a strainer, a pump, two flexible connectors, a check valve (non-return valve) and a riser pipe. For three consecutive years (i.e. 2005-2007), failure records of these components including emergency breakdowns and repair times were collected from both of the fresh and flushing water supply systems servicing the buildings. No water pump preventive maintenance was performed in that period.

3. Failure records

The total yearly failure counts logged and the restoration times estimated by the component repair engineers are shown in Table 2. For pump failures, two modes are presented – an overhaul needed and a replacement needed. The observed failure rate λ (h^{-1}) can be determined for a component using

M, the total number of installation components of the same type and N_f , the total failure counts recorded in the service period τ (=8760 h for one year),

$$\lambda = \frac{N_f}{M\tau} \quad \dots (1)$$

Except for those components that reported zero failures (i.e. the gate valves and strainers in the fresh water system), the geometric mean (GM) and geometric standard deviation (GSD) of λ obtained from the yearly records give the failure rate of a component as follows, where n is the number of yearly records available and $i=1\dots n$,

$$\ln GM = \frac{1}{n} \sum_{i=1}^n \ln \lambda_i \quad ; \quad \ln GSD = \sqrt{\frac{\sum_{i=1}^n (\ln \lambda_i - \ln GM)^2}{n}} \quad \dots (2)$$

According to the literature, water distribution component failure is rarely expected to occur (usually once in $>10^4$ hours of operation).⁵ Apparently, the zero-failure data indicated that a 26280-hour (=3-year) survey was insufficient for accurate analysis. Nevertheless, such data could be employed to improve the state of knowledge of the component failure rate via Bayes' theorem.⁶

Bayes' theorem relates the conditional and marginal/prior probabilities of events A and B, where B has a non-vanishing probability. Its key idea is that the probability of an event A given an event B depends not only on the relationship between events A and B but also on the marginal probability of

occurrence of each event. For example, if the failure rate of a component determined by a sample test is known to be 99% accurate, it could be due to 1% incorrect identification by the test (false positives), 1% missed cases (false negatives), or a mix of both. The application of Bayes' theorem allows calculations of the conditional probability of component failure, given an observed failure rate, for any of these three cases.

Hence, for A_i , a set of mutually exclusive and exhaustive events A describing the existing understanding of failure rates i (for $i=1,2,\dots$) of a specific component, given event B , a new failure-free observation for that component, the posterior probability $P(A_i|B)$ is defined as,

$$P(A_i | B) = \frac{P(A_i)P(B | A_i)}{\sum P(A_i)P(B | A_i)} \quad \dots (3)$$

$P(B|A_i)$ can be worked out as follows, where τ is the failure-free service period,

$$P(B | A_i) = [1 - P(A_i)]^\tau \quad \dots (4)$$

As it was difficult to obtain enough information to determine the precise shape of the failure distribution for the component concerned, a uniform prior probability distribution was assumed for the Bayesian analysis.⁷ Besides, Simpson's rule was applied to a normally distributed parameter ζ (or its transformation) with mean μ and standard deviation σ as expressed below⁸, where \min , pro and \max denotes the minimum, probable and

maximum values of the parameter ζ ,

$$\zeta \in \zeta(\mu, \sigma); \mu = \frac{1}{3}(\zeta_{\min} + \zeta_{\text{pro}} + \zeta_{\max});$$

$$\sigma = \frac{1}{18} \left\{ \left[\zeta_{\text{pro}} - \frac{1}{2}(\zeta_{\min} + \zeta_{\max}) \right]^2 + \frac{3}{4}(\zeta_{\max} - \zeta_{\min})^2 \right\} \quad \dots (5)$$

A prior distribution characterized by the geometric mean= 1.2×10^{-6} and geometric standard deviation=35, i.e. $\lambda_a \sim G(1.2 \times 10^{-6}, 35)$ as graphed in Figure 2, was suggested in this study. It was decided based on the failure probability range (0.1×10^{-6} - 15×10^{-6}) reported earlier for other water systems.⁵ Values of $P(B|A_i)$ for the gate valves and strainers where zero failures were observed were determined using Equation (4). Figure 2 also shows the posterior probabilities of these two components given by Equation (3) as the best estimates of the failure rates λ_a . The estimates were $0.02 \times 10^{-6} \text{ h}^{-1}$ and $0.04 \times 10^{-6} \text{ h}^{-1}$ for the gate valves and strainers respectively.

4. Simulations for system failures

Based on the component failure rates summarized in Table 2, simulations were carried out in order that the failure patterns of both the fresh and flushing water supply systems could be studied. Operation of a pump installation at any time was represented by either 'state 0' – installation

available or ‘state 1’ – installation failure. Any faulty installation component would lead to state 1 and a downtime τ_1 (h) (i.e. a period of state 1) till the repair was completed. The time between failures was defined as τ_0 (h) (i.e. a period of state 0). Hence, the entire operation period of a pump installation τ_∞ (h) can be expressed by a sum of time series τ_i , where $i=1,2,\dots,N_o$ and N_o is the total number of time periods of faulty-available pairs for the installation,

$$\tau_\infty = \tau_{0,1} + \tau_{1,1} + \tau_{0,2} + \tau_{1,2} + \dots + \tau_{0,N_o} + \tau_{1,N_o} = \sum \tau_0 + \sum \tau_1 \quad \dots (6)$$

Installation availability α_0 and installation unavailability α_1 are then given by,

$$\alpha_0 = \frac{\sum \tau_0}{\tau_\infty}; \alpha_1 = \frac{\sum \tau_1}{\tau_\infty}; \alpha_0 + \alpha_1 = 1 \quad \dots (7)$$

By making use of the non-constant component failure rates from Table 2, operations of a system composed of N_c components connected in series can be approximated via Monte Carlo simulations.

The component operation state S (in each hour) is described by,

$$S = \begin{cases} 1 & ; \sum_{i=1}^{N_c} S_i \geq 1 \\ 0 & ; \sum_{i=1}^{N_c} S_i = 0 \end{cases} \quad \dots (8)$$

where S_i is the operation state of a component i ,

$$S_i = \begin{cases} 1 & ; \vartheta \leq \lambda_i \\ 0 & ; \vartheta > \lambda_i \end{cases} ; \lambda_i = \lambda_{i,x} \in \tilde{\lambda}_i ; \int_{-\infty}^{\lambda_{i,x}} \tilde{\lambda}_i d\lambda_i = x \quad \dots (9)$$

In Equation (9), $\tilde{\lambda}_i$ is a distribution function for the component failure rate λ_i at percentile $x=\vartheta$, and $\vartheta \in [0,1]$ is a random number taken from a pseudo random number set generated by the prime modulus multiplicative linear congruential generator.⁹ The pseudo set has been tested and applied in a number of engineering applications.^{10,11,12}

In this study, a 10-year (i.e. $\tau_{\infty}=87600$ h) operation of 200 water pump installations – 100 fresh and 100 flushing – was simulated. The choice of the simulation period should be long enough to tentatively include more than 5 expected failure counts per installation component. An expression of the mean time to failure τ_m (h) given by a constant hazard rate model with an assumed exponential time to failure for each component i can be employed to justify the choice,¹³

$$\tau_m = \frac{1}{\sum_{i=1}^{N_c} \lambda_i} \quad \dots (10)$$

As Equation (10) gave an estimated mean time to failure of 15316 h (1.7 years) for the fresh water pump installations or 7765 h (0.9 years) for the flushing ones, i.e. an equivalence of 6 or more expected failure counts within 10 years per installation component, the simulation period was considered satisfactory.

Limited by the computer memory, the number of simulations required had to be balanced by accuracy and simulation run time. Parameter sensitivity to the number of simulations was examined by doubling the number of installations from 100 to 200 and the simulation time from 50 to 100 h per case. As the average values of the parameters concerned (such as τ_0 and τ_1) did not change significantly ($p > 0.4$, t-test), the simulations of 100 installations for each water supply system were selected for demonstrations.

For simplicity, some convenient parametric distributions, e.g. exponential, normal, lognormal and Gamma distributions, were chosen to fit the simulated data sets of downtime and time between failures in this pilot study. It was found that an exponential probability distribution, which is typical for many electronic components and complex industrial plants,¹⁴ would best describe the equipment failure pattern. Moreover, normal distribution (that is supposed to arise from additive effects of a large number of independent causative random variables) was used to approximate the random outcome.

The counts of τ_0 (h) and τ_1 (h) were obtained from the simulated time series of S. Figure 3 exhibits the counts of τ_0 (h) for both of the fresh and flushing water pump installations. τ_0 (h) can be described by exponential distributions ($p \geq 0.5$, Chi-square test) with a density function stated below, where the regression constant v_0 is 377 (or 0.46 with respect to percentage frequency) for the fresh water pump installations and 2641 (or 1.43 with respect to percentage frequency) for the flushing water pump installations,

$$\tilde{\tau}_0 \sim v_0 e^{-v\tau_0} \quad \dots (11)$$

The explanatory parameter values reported for the fresh and flushing water pump installations were $v=92 \times 10^{-6}$ and 247×10^{-6} corresponding to the mean times to failure $\tau_m=10817$ h (1.24 years) and 4048 h (0.46 years) respectively, i.e. about 30% and 50% lower than those estimated using the constant hazard rate model.

Figure 4 presents the downtimes τ_1 (h) for the fresh and flushing water pump installations in terms of frequency and percentage frequency. The failure counts observed in the flushing system were about 3 times higher and that was consistent with the shorter estimates of mean time to failure found for the system. Most failures (73% and 82% for the fresh and flushing water pump installations respectively) were in the downtime range from 1 to 5 hours; the rest (i.e. 27% and 18% respectively) had downtimes ≥ 13 hours. Over the 10-year simulation period, the total downtime for the 100 fresh water pump installations was 8990 hours with 709 failures and that for the 100 flushing water pump installations was 18988 hours with 2081 failures.

Figure 5 illustrates the downtimes for 100 fresh water supply systems and 100 flushing water supply systems which could be approximated by normal distributions ($p \geq 0.95$, Shapiro-Wilk's test). The average downtime and the corresponding availability for the fresh water supply system were 90 h (SD=63 h) and 0.90 while those for the flushing system were 190 h (SD=75

h) and 0.78 respectively.

To reduce installation unavailability, pumps are typically operated in parallel (i.e. 2 installations connected in parallel) as a means of flow control and for emergency back-up. The availability of a parallel pump installation $\alpha_{0,2L}$ is calculated using the following equation, where $\alpha_{0,L}$ is the availability of a pump installation,¹⁵ and the downtime τ_1 of each installation can be sampled from the parametric distributions shown in Figure 5 via Monte Carol simulation,

$$\alpha_{0,2L} = 1 - (1 - \alpha_{0,L})^2 ; \alpha_{0,L} = 1 - \frac{\tau_1}{\tau_\infty} ; \tau_1 \in \tilde{\tau}_1 \quad \dots (12)$$

The total downtime of a parallel pump installation τ_{2L} (h) in an operation period τ_∞ (h) is,

$$\tau_{2L} = \tau_\infty (1 - \alpha_{0,2L}) \quad \dots (13)$$

It can be approximated by a Gamma distribution ($p > 0.95$, Chi-square test), $\tilde{\tau}_{2L} \sim G_m(\beta_1, \beta_2)$, for describing the performance in terms of the allowable downtime of the water pump installations. The probability of a parallel pump installation P_{2L} associated with a total downtime $\tau_{2L} \geq \tau_{2L}^*$ (h) in τ_∞ (h) can be determined by,

$$P_{2L}(\tau_{2L} \geq \tau_{2L}^*) = 1 - \int_0^{\tau_{2L}^*} G_m(\beta_1, \beta_2) d\tau_{2L} \quad \dots (14)$$

Figure 6 shows the total downtime distribution of parallel water pump installations. If the maximum allowable downtime per year=1 h, then P_{2L} would be 40×10^{-6} and 0.01 respectively for the fresh and flushing water pump installations in Hong Kong residential buildings.

5. Conclusion

Currently in Hong Kong, systematic considerations for the influence of various installation arrangements and usage patterns on the reliability of a water supply system are lacking. Together with some reliability data published elsewhere, this study applied Bayesian analysis to improve our understanding of the downtime characteristics of water pump installations. Based on a survey on water pump failures conducted in Hong Kong from 2005 to 2007 for typical high-rise residential buildings, the failure patterns of installation components were estimated via Monte Carlo simulations. From the virtual operations of 100 parallel pump sets over a period of 10 years, the mean time to failure, total downtime, failure counts and system availability for the fresh water pump installations were 1.24 years, 8990 hours, 709 and 90%, while those for the flushing water pump installations were 0.46 years, 18988 hours, 2081 and 78%, respectively. It is noted that for Hong Kong, the seawater flushing water closet is considered as a cost-effective and environmentally friendly water conservation measure. This study reported that seawater flushing water pumps were associated with

more failures but shorter downtimes than the fresh water supply pumps. The results would be useful in the calculation of water supply availability for high-rise residential buildings while keeping the balance between maintenance cost and system reliability. This study also demonstrated a method for reliability modelling of water supply for high-rise residential buildings.

Acknowledgement

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References

1. Lee CK, Yu CW. Conservation of water resources - use of sea water for flushing in Hong Kong. *Journal of water supply research and technology-Aqua* 1997;46(4):202-9.
2. Li XZ, Luk SF, Tang SL. Sustainability of toilet flushing water supply in Hong Kong. *Water and Environment Journal* 2005;19(2):85-90.
3. Tao Z. Application of GIS and CARE-W Systems on Water Distribution Networks, Skärholmen, Stockholm, Sweden. MSc Thesis, 781027-A242, International Master Programme of Environmental Engineering &

Sustainable Infrastructure, Kungliga Tekniska Högskolan (Royal Institute of Technology), Sweden; 2006.

4. Wong LT, Mui KW. An occupant load survey for residential buildings in Hong Kong. *International Journal for Housing Science and its applications* 2006;30(3):195-204.

5. Fong NK. Reliability study on sprinkler system to be installed in old high-rise buildings. *International Journal on engineering Performance-Based Fire Codes* 2000;2(2):61-7.

6. Gelman A, Carlin JB, Stern HS, Rubin DB. *Bayesian data analysis*, 2nd Ed. Boca Raton: Chapman & Hall/CRC; 2003.

7. Aitken CGG, Taroni F. *Statistics and the evaluation of evidence for forensic scientists*, 2nd Ed. England: John Wiley & Sons; 2004.

8. Lerche I, Paleologos EK. *Environmental risk analysis*, New York: McGraw-Hill; 2001.

9. Park SK, Miller KW. Random number generators: good ones are hard to find. *Communications of the ACM* 1988;31(10):1192-201.

10. Wong LT, Mui KW. A transient ventilation demand model for air-conditioned offices. *Applied Energy* 2008;85(7):545-54.

11. Wong LT, Mui KW. Stochastic modelling of water demand by domestic washrooms in residential tower blocks. *Water and Environment Journal* 2008;22(2):125-30.

12. Wong LT, Mui KW, Guan Y. Shower water heat recovery in high-rise residential buildings of Hong Kong. *Applied Energy* 2010;87(2):703-9.
13. Modarres M, Joglar-Billoch F. Reliability, In *The SFPE handbook of fire protection engineering*, 3rd Edition, Chapter 5-3, Maryland, USA: Society of Fire Protection Engineers; 2002.
14. Jardine AKS, Tsang AHC. *Maintenance, replacement, and reliability*. London: CRC Press; 2006.
15. Qi Y. A study on the reliability of fire water supply system in high-rise buildings. *Fire Technology* 2002;38(1):71-9.

Table 1: Data of the 46 high-rise residential buildings surveyed

Residential building type	Sample size (No.)	Number of apartments	Floor area (m ²)	Year of completion	Number of installation components for fresh water system and flushing water system					
					Gate valve	Check valve	Flexible connector	Strainer	Pump	Riser pipe
(1) Public	12	18312	18-68	1987-1989	48	24	48	24	24	12
(2) Subsidized	16	33022	42-76	1983-1990	64	32	64	32	32	16
(3) Private	18	2022	48-89	1989-1992	72	36	72	36	36	18

Table 2: Failure records of components in water pump installations

Component	Restoration time (h)	Yearly failure counts N_f for building types (1), (2) and (3) ^b									Observed failure rate λ for three consecutive years ($\times 10^{-6} \text{ h}^{-1}$)					Reference failure rate ⁵ λ_0 ($\times 10^{-6} \text{ h}^{-1}$)	Component failure rate λ_1 ($\times 10^{-6} \text{ h}^{-1}$)	
		(1)			(2)			(3)			(1)	(2)	(3)	GM	GSD		GM	GSD
		(1)	(1)	(1)	(2)	(2)	(2)	(3)	(3)	(3)	(1)	(2)	(3)	GM	GSD		(1)	(2)
<i>Fresh water system</i>																		
Gate valve	1-3	0	0	0	0	0	0	0	0	0	0	0	0	--	--	0.1-50	0.02	1.28
Check valve	1-3	0	1	0	2	0	2	1	2	1	4.8	7.1	4.2	5.0	1.5	1-8	5	1.5
Flexible connector	1-3	1	0	1	3	3	9	5	3	2	1.6	8.9	5.3	4.9	1.9	--	4.9	1.9
Strainer	1-3	0	0	0	0	0	0	0	0	0	0	0	0	--	--	--	0.04	2.38
Pump ^a	24-72	2	4	4	5	7	3	2	5	2	16	18	9.5	13	1.7	0.7-200	13	1.7
	(12-48)	(1)	(0)	(0)	(2)	(0)	(0)	(0)	(1)	(0)	(4.8)	(7.1)	(3.2)	(4.8)	(1.5)		(4.8)	(1.5)
Riser pipe	1-3	3	3	4	7	2	9	2	15	4	32	43	44	33	1.9	0.2	33	1.9
<i>Flushing water system</i>																		

Gate valve	1-3	4	4	3	0	9	13	8	2	1	8.7	20	5.8	7.9	2.4	0.1-50	7.9	2.4
Check valve	1-3	0	4	3	5	9	5	7	6	3	17	23	17	18	1.4	1-8	18	1.4
Flexible connector	1-3	2	10	5	0	2	1	0	1	1	13	2.7	1.6	4.2	2.9	--	4.2	2.9
Strainer	1-3	1	1	1	0	5	3	1	1	0	4.8	14	3.2	5.7	1.9	--	5.7	1.9
Pump ^a	24-72 (12-48)	6 (1)	7 (2)	5 (4)	3 (5)	1 (11)	5 (7)	9 (3)	3 (13)	9 (5)	29 (11)	11 (27)	22 (22)	17 (17)	2.1 (2.0)	0.7-200	17 (17)	2.1 (2.0)
Riser pipe	1-3	2	22	88	1	2	5	4	12	8	355	19	51	47	4.3	0.2	47	4.3

^a an overhaul needed, and a replacement needed in brackets; ^b building types (1), (2) and (3) are public, subsidized and private housing estates respectively; -- = not applicable;

Figure 1: A typical water pump installation for high-rise residential buildings

Figure 2: Failure rates of gate valves and strainers for fresh water pump installations

Figure 3: Times between failures for fresh and flushing water pump installations

Figure 4: Downtimes of fresh and flushing water pump installations

Figure 5: Downtime of water pump installations

Figure 6: Total downtimes of parallel water pump installations

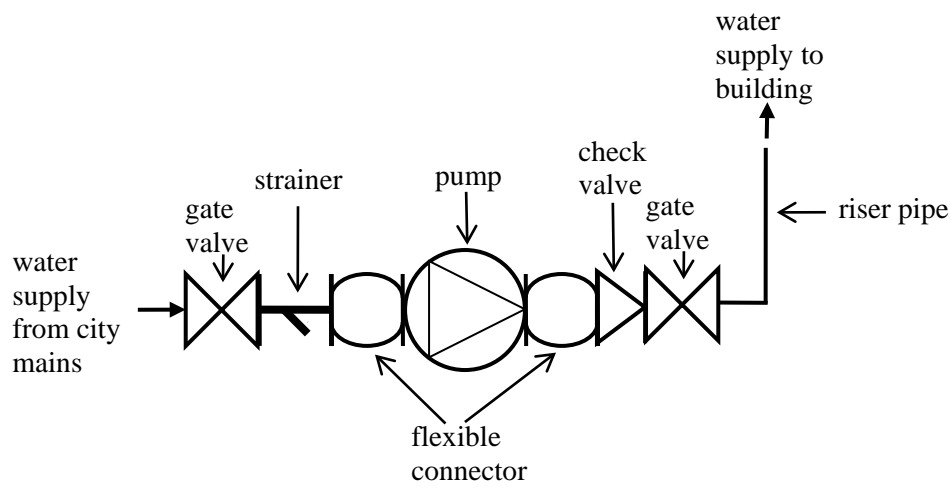
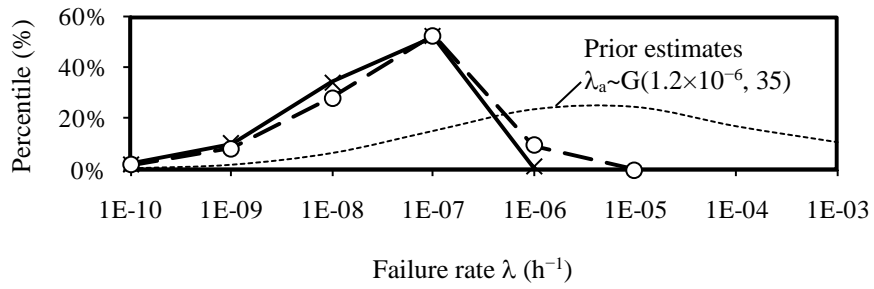


Figure 1: A typical water pump installation for high-rise residential buildings



Posterior estimates for fresh water pump installations

- Gate valve $\lambda_{a'} \sim G(0.02 \times 10^{-6}, 1.28)$
- - Strainer $\lambda_{a'} \sim G(0.04 \times 10^{-6}, 2.38)$

Figure 2: Failure rates of gate valves and strainers for fresh water pump installations

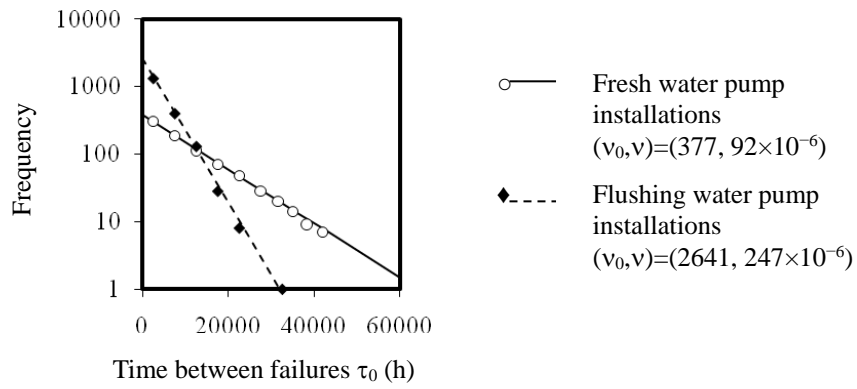
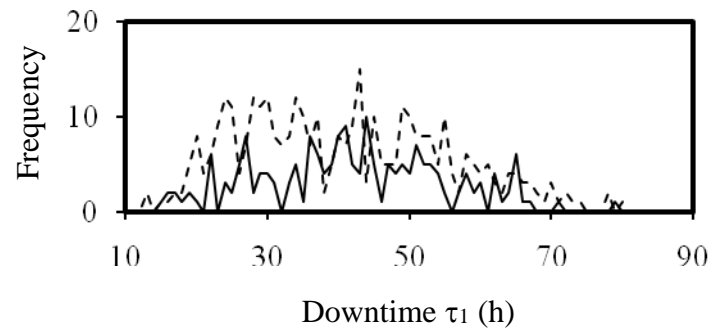
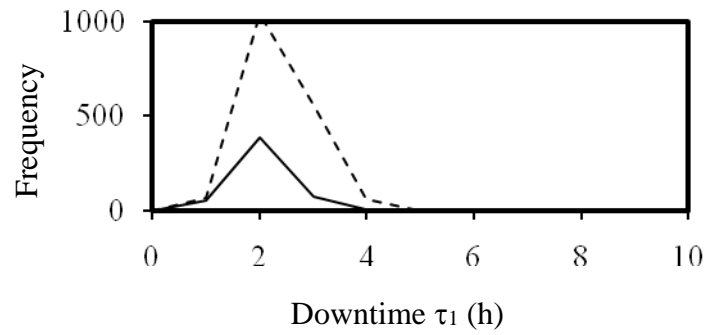
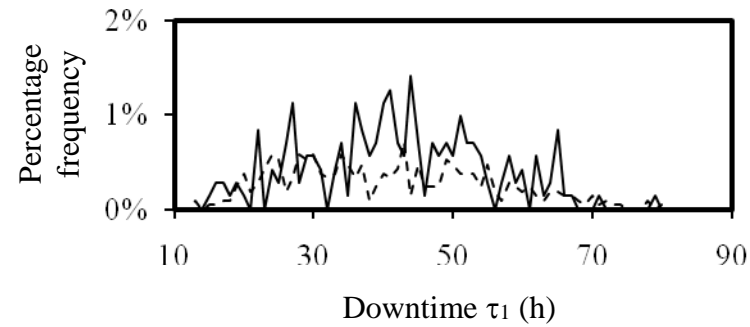
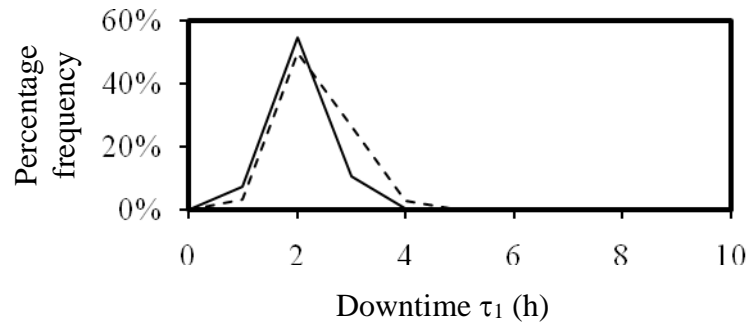


Figure 3: Times between failures for fresh and flushing water pump installations



(a) Frequency



(b) Percentage frequency

— Fresh water pump installations
 ---- Flushing water pump installations

Figure 4: Downtimes of fresh and flushing water pump installations

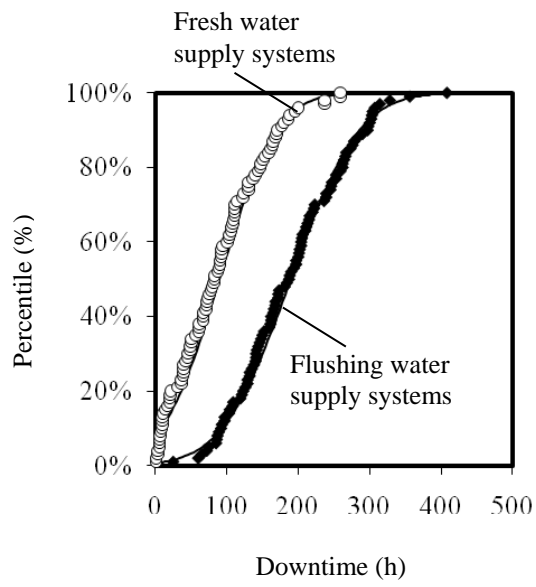


Figure 5: Downtime of water pump installations

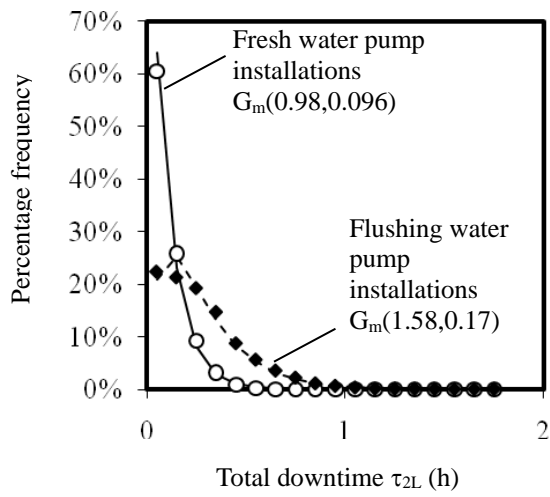


Figure 6: Total downtimes of parallel water pump installations